5.4 Joint distributions and independence (discrete) we into now, only

Most applications of probability to engineering statistics involve not one but several random variables. In some cases, the application is intrinsically multivariate.

Example 5.32. Consider the assembly of a ring bearing with nominal inside diameter 1.00 in. on a rod with nominal diameter .99 in. If

X =the ring bearing inside diameter

Y =the rod diameter

One might be interested in

Even when a situation is univariate, samples larger than size 1 are essentially always used in engineering applications. The n data values in a sample are usually thought of as subject to chance and their simultaneous behavior must then be modeled.

This is actually a very broad and difficult subject, we will only cover a brief introduction to the topic: **jointly discrete random variables**.

5.4.1 Joint distributions

For several discrete random variable, the device typically used to specify probabilities is a *joint probability function*. The two-variable version of this is defined.

Definition 5.21. A joint probability function (joint pmf) for discrete random variables X and Y is a nonnegative function f(x, y), giving the probability that (simultaneously) X takes the values x and Y takes the values y. That is,

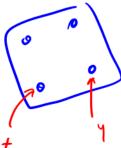
$$f(x,y) = P[X = x \text{ and } Y = y]$$

Properties:

$$\sum_{x,y} f(x,y) = 1$$

For the discrete case, it is useful to give f(x, y) in a **table**.

Example 5.33 (Two bolt torques, cont'd). Recall the example of measure the bolt torques on the face plates of a heavy equipment component to the nearest integer. With



• † Y	,	X = the next torque recorded for bolt 3 $Y = $ the next torque recorded for bolt 4 $P[X = X, Y = Y]$										d to eger
the joint	proba	_	tunctio	on, $f(x)$,y), is							
	y\x	11	12	13	14	15	16	17	18	19	20	
7	20	0	0	0	0	0	0	0	2/34	2/34	1/34	
↓	19	0	0	0	(0)	0	0	2/34	0	0	0	
	18	0	0	1/34	1/34	0	0	1/34	1/34	1/34	0	
	17	0	0	0	0	2/34	1/34	1/34	2/34	0	0	
	16	0	0	0	1/34	2/34	2/34	0	Û	2/34	0	
	15	1/34	1/34	0	0	3/34	0	0	0	0	0	
	14	0	0	0	0	1/34	0	0	2/34	0	0	
_	13	0	0	0	0	1/34	0	0	b	0	0	
												•

$$P[X = 18 \text{ and } Y = 17] = \frac{2}{34}$$

$$P[X = 14 \text{ and } Y = 19] = 0$$

By summing up certain values of f(x, y), probabilities associated with X and Y with patterns of interest can be obtained.

Consider:

X

P(polt 3 requires more) $P[X \ge Y] = \rho(x=13, Y=13) + \rho(x=14, Y=13) + \rho(x=14, Y=14) + \dots$ $= f(13,13) + f(14,13) + \dots + f(26,13) + \rho(x=20, Y=13)$ $y \setminus x$ × ×

× × X × × ×

P[X = 17] = f(17,20) + ... + f(17,13) $y \setminus x$

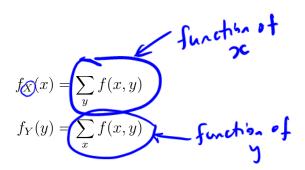
X

5.4.2 Maginal distributions

Jecause we put the result in factures of the table (really).

In a bivariate problem, once can add down columns in the (two-way) table of f(x, y) to get values for the probability function of X, $f_X(x)$ and across rows in the same table to get values for the probability distribution of Y, $f_Y(\dot{y})$.

Definition 5.22. The individual probability functions for discrete random variables X and Y with joint probability function f(x,y) are called marginal probability functions. They are obtained by summing f(x,y) values over all possible values of the other variable.



Example 5.34 (Torques, cont'd). Find the marginal probability functions for X and Y from the following joint pmf.

y\x	11	12	13	14	15	16	17	18	19	20	fy(9)
20	0	0	0	0	0	0	0	2/34	2/34	1/34	5/34
19	0	0	0	0	0	0	2/34	0	0	0	2/34
18	0	0	1/34	1/34	0	0	1/34	1/34	1/34	0	5131
17	0	0	0	0	2/34	1/34	1/34	2/34	0	0	6/34
16	0	0	0	1/34	2/34	2/34	0	0	2/34	0	7/31
15	1/34	1/34	0	0	3/34	0	0	0	0	0	5/34
14	0	0	0	0	1/34	0	0	2/34	0	0	3/31
13	0	0	0	0	1/34	0	0	0	0	0	1/31
£ (xc)	<u></u>	1	1	2	1	3	4	7	5		ľ
Xcas	34	34	34	34	74	34	34	34	34	34	l

Getting marginal probability functions from joint probability functions begs the question whether the process can be reversed. Can we find joint probability functions from marginal probability functions? No! (sometimes yes, more later) Consider X and Y with joint distributions

5.4.3Conditional distributions

When working with several random variables, it is often useful to think about what is expected of one of the variables, given the values assumed by all others.

For example, in the bolt torque example, you ought to have some expectations for bolt 4 torque if you know bolt 3 took 15 ft 15 to lessen.

Definition 5.23. For discrete random variables X and Y with joint probability function f(x,y), the conditional probability function of X given Y=y is the function of x

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{f(x,y)}{\sum\limits_{x} f(x,y)} = \frac{f(x,y)}{\text{marginal of } Y}$$

XIY " X given Y"

and the conditional probability function of Y given X = x is the function of y

$$f_{Y|X}(y|x) = \frac{f(x,y)}{f_X(x)} = \frac{f(x,y)}{\sum\limits_{y} f(x,y)} = \frac{\text{joint}}{\text{maginal of } X}$$

Example 5.35 (Torque, cont'd). For the torque example with the following joint distribution, find the following:

1.
$$f_{Y|X}(20|18) = P(Y = 20 \text{ given } X = 18) = \frac{f(18, 20)}{f_{X}(18)} = \frac{2/34}{7/34} = \frac{2}{7}$$

- 2. $f_{Y|X}(y|15)$
- 3. $f_{Y|X}(y|20)$
- 4. $f_{X|Y}(x|18)$

			•		\sim						
$y \setminus x$	11	12	13	14	/15	16	17	18	19	20	$f_Y(y)$
20	0/34	0/34	0/34	0/34	0/34	0/34	0/34	2/34	2/34	1/34	5/34
19	0/34	0/34	0/34	0/34	0/34	0/34	2/34	0/34	0/34	0/34	2/34
→ 18	0/34	0/34	1/34	1/34	0/34	0/34	1/34	1/34	1/34	0/34	5/34
17	0/34	0/34	0/34	0/34	2/34	1/34	1/34	2/34	0/34	0/34	6/34
16	0/34	0/34	0/34	1/34	2/34	2/34	0/34	0/34	2/34	0/34	7/34
15	1/34	1/34	0/34	0/34	3/34	0/34	0/34	0/34	0/34	0/34	5/34
14	0/34	0/34	0/34	0/34	1/34	0/34	0/34	2/34	0/34	0/34	3/34
13	0/34	0/34	0/34	0/34	1/34	0/34	0/34	0/34	0/34	0/34	1/34
$f_X(x)$	1/34	1/34	1/34	2/34	9/34	3/34	4/34	(7/34)	5/34	1/34	34/34

2.
$$f_{y|x}(y|15) = \frac{f(|5,y)}{f_{x}(15)}$$

3.
$$f_{Y|X}(y|20) = \frac{f(20, y)}{f_{x}(20)}$$

4.
$$f_{x|y}(x|18) = \frac{f(x,18)}{f_y(18)}$$

2.
$$f_{y|x}(y|15) = \frac{f(|5,y)}{f_{x}(15)}$$
 $g = \frac{f_{y|x}(y|15)}{f_{y|x}(y|15)}$
 $g = \frac{f_{y|x}(y|15)}{f_{x}(15)}$
 $g = \frac{f_{y|x}(y|15)}{f_{y|x}(y|15)}$
 $g = \frac{f_{y|x}(y|15)}{g_{y|x}(y|15)}$
 $g = \frac{f_{y|x}(y$

$$\begin{array}{c|cccc} X & f_{X|Y}(x|Y=18) \\ \hline 11 & O \\ 12 & O \\ 13 & y_{3Y}/(5/34) = \frac{1}{5} \\ 14 & y_{3Y}/(5/34) = \frac{1}{5} \\ 15 & O \\ 16 & O \\ 17 & y_{3Y}/(5/34) = \frac{1}{5} \\ 18 & y_{3Y}/(5/34) = \frac{1}{5} \\ 19 & y_{3Y}/(5/34) = \frac{1}{5} \\ 19 & y_{3Y}/(5/34) = \frac{1}{5} \\ 20 & O \\ \end{array}$$

5.4.4 Independence

Recall the following joint distribution:

y x	1	2	3	$f_Y(y)$
3	0.08	0.08	0.04	0.20
2	0.16	0.16	0.08	0.40
1	0.16	0.16	0.08	0.40
$f_X(x)$	0.40	0.40	0.20	1.00

What do you notice?

Each
$$P[x=x, Y=y] = P[x=x] P[y=y]$$

Also, $f_{y|x}(y|3) = \frac{f(3,y)}{f_{x}(5)} \Rightarrow \frac{y | f_{y|x}(y|3)}{1 | .08/.2 = .7}$

So, fylx (y13) = fy(y). Actually, this is the for all of

i.e. Knowing what value X takes, doesn't malk in questions about Y!

Definition 5.24. Discrete random variables X and Y are <u>independent</u> if their joint distribution function f(x,y) is the product of their respective marginal probability functions. This is, independence means that

$$\begin{split} \mathbf{P} \big[\mathbf{X} = \mathbf{x}, \mathbf{Y} = \mathbf{y} \big] = f(x, y) &= f_X(x) f_Y(y) & \text{for all } x, y. \\ &= \mathbf{P} \big[\mathbf{X} = \mathbf{x} \big] \big[\mathbf{y} = \mathbf{y} \big] & \text{(for discrete)} \end{split}$$

If this does not hold, then X and Y are dependent.

Alternatively, discrete random variables X and Y are independent if for all x and y,

If X and Y are not only independent but also have the same marginal distribution, then they are independent and identically distributed (iid).

5.5 Functions of several random variables

joint distributions

We've now talked about ways to simultaneously model several random variables. An important engineering use of that material is in the analysis of system output that are functions of random inputs.

5.5.1 Linear combinations

For engineering purposes, it often suffices to know the mean and variance for a function of several random variables, $U = g(X_1, X_2, ..., X_n)$ (as opposed to knowing the whole distribution of U). When g is **linear**, there are explicit functions.

Proposition 5.1. If $X_1, X_2, ..., X_n$ are n independent random variables and $a_0, a_1, ..., a_n$ are n + 1 constants, then the random variable $U = a_0 + a_1X_1 + a_2X_2 + \cdots + a_nX_n$ has mean

$$EU=a_0+a_1EX_1+a_2EX_2+\cdots+a_nEX_m$$
 — this holds even if X_1,\dots,X_n are not

and variance

$$VarU = a_1^2 (X_1 + a_2^2 (X_2 + \dots + a_n^2) (X_3)$$

idea of proof:

For n=2 joint probability fundom
$$f(x_1,x_2) = P[x_1=x_1, X_2=x_2]$$
.
Define $U = a_0 + a_1x_1 + a_2x_2$
 $EU = E[a_0 + a_1x_1 + a_2x_2]$

$$= \sum_{x_1,x_2} (a_0 + a_1 x_1 + a_2 x_2) f(x_1,x_2)$$

$$= \sum_{x_1,x_2} \sum_{\alpha_0} a_0 f(x_0,x_2) + \sum_{\alpha_1,\alpha_2} \sum_{\alpha_1,\alpha_2} a_1 f(x_1,x_2) + \sum_{\alpha_1,\alpha_2} \sum_{\alpha_2,\alpha_2} a_2 f(x_1,x_2)$$

$$= a_0 \left(\sum_{x_1, x_2} \int_{x_1, x_2} \int_{x_1} \int_{x_2} \int$$

$$= a_0 + a_1 \underbrace{\sum_{x_1} x_1 f_{x_1}(x_1)}_{61} + a_2 \underbrace{\sum_{x_2} f_{x_2} f_{x_2}}_{62}$$

$$\times_2 = a_0 + a_1 E \times_1 + a_2 E \times_2$$

Check on your own, same ideas hold for Varu = a,2 Varx, + a2 VarX2.//

Example 5.36. Say we have two independent random variables X and Y with EX = 3.3, VarX = 1.91, EY = 25, and VarY = 65. Find the mean and variance for

$$U = 3 + 2X - 3Y$$

$$V = -4X + 3Y$$

$$W = 2X - 5Y$$

$$Z = -4X - 6Y$$

$$EU = E(3+2x-3y)$$

$$= \frac{3}{3} + 2Ex - 3Ey$$

$$= \frac{3}{3} + 2(3.5) - 3(25) = -65.4$$

$$EZ = E[-4x - 6y]$$

= $-4EX - 6EY$
= $-4(3.3) - 6(25) = -163.2$

$$V \sim M = V \sim (3 + 2 \times - 3 Y)$$

= $2^2 V \sim X + (-3)^2 V \sim Y$
= $4 (1.91) + 9 (65) = 592.64$

$$VarV = Var(-4x+3y)$$

= $(-4)^2 Var x + 3^2 Var y$
= $16(1.91) + 9(65) = 615.56$

$$V_{WW} = V_{W} (2x - 5y)$$

 $= 2^2 V_{W} x + (-5)^2 V_{W}$
 $= 4(1.91) + 25(65) = 1632.64$

$$Var = Var(-4x - 6y)$$

$$= (-4)^{2} Var x + (-6)^{2} Var y$$

$$= 16(1.91) + 36(65) =$$

$$= 2370.56$$

and X and Y are in dependent.

Example 5.37. Say $X \sim Binomial(n = 10, p = 0.5)$ and $Y \sim Poisson(\lambda = 3)$. Calculate the mean and variance of Z = 5 + 2X - 7Y.

First note

$$EX = n \cdot p = 10 \cdot 0.5 = 5$$

 $Var X = n \cdot p (1 - p) = 10 \cdot 0.5 \cdot 0.5 = 2.5$
 $EY = \lambda = 3$
 $Var Y = \lambda = 3$

$$TL_{\Lambda}$$
 EZ = E(5+2X-7Y)
= 5+ 2EX -7EY
= 5+2(5) -7(3)
= -6

$$Var = Var (5 + 2x - 7Y)$$

= $2^2 Var X + (-7)^2 Var Y$.
= $4(2.5) + 49(3)$
= 157

independent the little and a find the stand

A particularly important use of Proposition 5.1 concerns n (iid) random variables where each $a_i = \frac{1}{n_i}$ for i = 1, ..., n

X17-17 Xn are conceptually equivalent to wardon selectors (with replacement) from a single numerical population.

We can find the mean and variance of the random variable

sarple mean.
$$\overline{X} = \overline{n} X_1 + \cdots + \frac{1}{n} X_n = \frac{1}{n} \sum_{i=1}^n X_i$$

as they relate to the population parameters $\mu = EX_i$ and $\sigma^2 = VarX_i$.

X17-7, Xn are identically Mistributed, EXF...= EXn=M VarXi=--= VarXn=62

For independent variables X_1, \ldots, X_n with common mean μ and variance σ^2 ,

$$E\overline{X} = E\left[\frac{1}{n}X_{1} + \dots + \frac{1}{n}X_{n}\right]$$

$$= \frac{1}{n}EX_{1} + \dots + \frac{1}{n}EX_{n} \quad (\text{prop 5.1})$$

$$= \frac{1}{n}M + \dots + \frac{1}{n}M$$

So, The expected value of the sample mean is population mean.

$$Var \overline{X} = Var \left[\frac{1}{n}X_1 + ... + \frac{1}{n}X_n\right]$$

$$= \left(\frac{1}{n}\right)^2 Var X_1 + ... + \left(\frac{1}{n}\right)^2 Var X_n \qquad (par 5.1)$$

$$= \frac{1}{n^2} 6^2 + ... + \frac{1}{n^2} 6^2$$

$$= \frac{1}{n^2} \cdot n \cdot 6^2 = 6^2 / n$$

The variance of the sample mean for a sample of size in is the population variance divided by the sample size.

ie - as the sample size grows, the variability of the sample mean decreases.

Example 5.38 (Seed lengths). One botanist measured the length of 10 seeds from the same plant. The seed lengths measurements are X_1, X_2, \ldots, X_{10} . Suppose it is known that the seed lengths are iid with mean $\mu = 5$ mm and variance $\sigma^2 = 2$ mm.

Calculate the mean and variance of the average of 10 seed measurements.

$$\overline{X}$$
 = average of 10 measurements
= $\frac{1}{10} \sum_{i=1}^{\infty} X_i$
Since X_i and with $\mu = 5$, $6^2 = 2$,
 $E\overline{X} = \mu = 5$ mm
Var $\overline{X} = \frac{1}{10}6^2 = \frac{2}{10} = 0.2$

5.5.2 Central limit theorem

One of the most frequently used statistics in engineering applications is the sample mean. We can relate the mean and variance of the probability distribution of the sample mean to those of a single observation when an iid model is appropriate.

Proposition 5.2. If X_1, \ldots, X_n are iid random variable (with mean μ and variance σ^2), then for large n, the variable \overline{X} is approximately normally distributed. That is,

$$\overline{X} \sim Normal\left(\mu, \frac{\sigma^2}{n}\right)$$

This is one of the **most important** results in statistics.

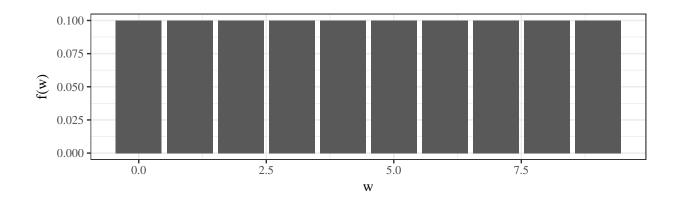
Example 5.39 (Tool serial numbers). Consider selecting the last digit of randomly selected serial numbers of pneumatic tools. Let

 W_1 = the last digit of the serial number observed next Monday at 9am

 W_2 = the last digit of the serial number observed the following Monday at 9am

A plausible model for the pair of random variables W_1, W_2 is that they are independent, each with the marginal probability function

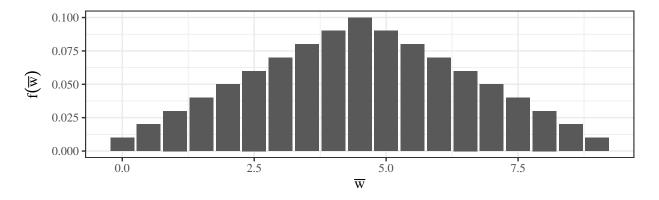
$$f(w) = \begin{cases} .1 & w = 0, 1, 2, \dots, 9 \\ 0 & \text{otherwise} \end{cases}$$



With EW = 4.5 and VarW = 8.25.

Using such a distribution, it is possible to see that $\overline{W} = \frac{1}{2}(W_1 + W_2)$ has probability distribution

	• (/		$f(\overline{w})$		• (/		• ()		• (/
0.00	0.01	2.00	0.05	4.00	0.09	6.00	0.07	8	0.03
0.50	0.02	2.50	0.06	4.50	0.10	6.50	0.06	8.5	0.02
			0.07					9	0.01
1.50	0.04	3.50	0.08	5.50	0.08	7.50	0.04		



Comparing the two distributions, it is clear that even for a completely flat/uniform distribution of W and a small sample size of n=2, the probability distribution of W looks more bell-shaped that the underlying distribution.

Now consider larger and larger sample sizes, $n=1,\ldots,40$:

Click for video...

Example 5.40 (Stamp sale time). Imagine you are a stamp salesperson (on eBay). Consider the time required to complete a stamp sale as S, and let

 $\overline{S}=$ the sample mean time required to complete the next 100 sales

Each individual sale time should have an $Exp(\alpha=16.5s)$ distribution. We want to consider approximating $P[\overline{S}>17]$.

Example 5.41 (Cars). Suppose a bunch of cars pass through certain stretch of road. Whenever a car comes, you look at your watch and record the time. Let X_i be the time (in minutes) between when the i^{th} car comes and the $(i+1)^{th}$ car comes for $i=1,\ldots,44$. Suppose you know the average time between cars is 1 minute. Find the probability that the average time gap between cars exceeds 1.5 minutes.

Example 5.42 (Baby food jars, cont'd). The process of filling food containers appears to have an inherent standard deviation of measured fill weights on the order of 1.6g. Suppose we want to calibrate the filling machine by setting an adjustment knob and filling a run of n jars. Their sample mean net contents will serve as an indication of the process mean fill level corresponding to that knob setting.

You want to choose a sample size, n, large enough that there is an 80

Example 5.43 (Printing mistakes). Suppose the number of printing mistakes on a page
follows some unknown distribution with a mean of 4 and a variance of 9. Assume that number of printing mistakes on a printed page are iid.
1. What is the approximate probability distribution of the average number of printing mistakes on 50 pages?
2. Can you find the probability that the number of printing mistakes on a single page is less than 3.8?
3. Can you find the probability that the average number of printing mistakes on 10 pages is less than 3.8?
4. Can you find the probability that the average number of printing mistakes on 50 pages is less than 3.8?